

# SECTION 4 – GLOBAL GAMES

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Gov 2006 – Harvard

February 18, 2020

# BAYESIAN UPDATING WITH CONTINUOUS DISTRIBUTIONS

I will tell you a trick that people doing Bayesian statistics use a lot.

Let's revisit something we did in class.

- There is a parameter  $\theta_0$  we don't know.
- We have a flat improper prior:  $p(\theta_0) = c > 0$ .
- We observe a signal  $\theta_1 | \theta_0 \sim N(\theta_0, \sigma^2)$ .
- We update our belief about  $\theta_0$ .

So, we want to know the distribution  $\theta_0 | \theta_1$ .

Bayes rule:

$$p(\theta_0 | \theta_1) = \frac{p(\theta_1 | \theta_0)p(\theta_0)}{p(\theta_1)}.$$

The numerator is easy: we are given  $p(\theta_1 | \theta_0)$  and  $p(\theta_0)$ .

The denominator is hard to calculate.

# THE TRICK

Key observation: **it's just a constant.**

Key observation 2: if you know a PDF  $p(x)$  up to a constant, then you know it, because the constant is determined by the formula  $\int p(x) dx = 1$ .

If you know that  $p(x) = kf(x)$  but you don't know  $k$ , you can calculate it:

$$1 = \int p(x) dx = \int kf(x) dx = k \int f(x) dx,$$

so

$$k = \frac{1}{\int f(x) dx}.$$

**We can ignore constant factors.**

Notation:  $p(x) \propto f(x)$  means that  $p(x) = kf(x)$ , where  $k > 0$  is a constant.

Bayes rule is  $p(\theta_0|\theta_1) = \frac{p(\theta_1|\theta_0)p(\theta_0)}{p(\theta_1)}$ , but, since the denominator is a constant, we can write it simply as

$$p(\theta_0|\theta_1) \propto p(\theta_1|\theta_0)p(\theta_0).$$

(If you read Bayesian stats papers, they do this all the time.)

Going back to our example, we have

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Going back to our example, we have

$$\begin{aligned} p(\theta_0|\theta_1) &\propto p(\theta_1|\theta_0)p(\theta_0) = \\ &= \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2\sigma^2}(\theta_1 - \theta_0)^2\right] c \propto \\ &\propto \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2\sigma^2}(\theta_0 - \theta_1)^2\right]. \end{aligned}$$

Now, this function of  $\theta_0$  integrates to 1, so it must be exactly the PDF. It is clearly the PDF of a normal  $N(\theta_1, \sigma^2)$ .

Hence,  $\theta_0|\theta_1 \sim N(\theta_1, \sigma^2)$ , which we knew already.

## ADDING A SECOND SIGNAL

Suppose that we observe another signal  $\theta_2|\theta_0 \sim N(\theta_0, \tau^2)$ , independent of  $\theta_1|\theta_0$ . We want the posterior  $\theta_0|\theta_1, \theta_2$ .

Bayes rule:

$$\begin{aligned} p(\theta_0|\theta_1, \theta_2) &\propto p(\theta_1, \theta_2|\theta_0)p(\theta_0) = \\ &= p(\theta_1|\theta_0)p(\theta_2|\theta_0)p(\theta_0) \propto \\ &\propto \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2\sigma^2}(\theta_1 - \theta_0)^2\right] \frac{1}{\sqrt{2\pi}\tau} \exp\left[-\frac{1}{2\tau^2}(\theta_2 - \theta_0)^2\right] \propto \\ &\propto \exp\left[-\frac{1}{2\sigma^2}(\theta_1 - \theta_0)^2 - \frac{1}{2\tau^2}(\theta_2 - \theta_0)^2\right] = \\ &= \exp\left[-\frac{1}{2\sigma^2}(\theta_0^2 - 2\theta_1\theta_0 + \theta_1^2) - \frac{1}{2\tau^2}(\theta_0^2 - 2\theta_2\theta_0 + \theta_2^2)\right] = \\ &= \exp\left[-\frac{1}{2}\left(\frac{1}{\sigma^2} + \frac{1}{\tau^2}\right)\theta_0^2 - \left(\frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2}\right)\theta_0 - \frac{\theta_1^2}{\sigma^2} - \frac{\theta_2^2}{\tau^2}\right] \propto \\ &\propto \exp\left[-\frac{1}{2}\left(\frac{1}{\sigma^2} + \frac{1}{\tau^2}\right)\theta_0^2 - \left(\frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2}\right)\theta_0\right] \end{aligned}$$

$$\begin{aligned}
p(\theta_0|\theta_1, \theta_2) &\propto \\
&\propto \exp \left[ -\frac{1}{2} \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) \theta_0^2 - \left( \frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2} \right) \theta_0 \right] = \\
&= \exp \left[ -\frac{1}{2} \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) \left( \theta_0 - \frac{\frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}} \right)^2 + \frac{\left( \frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2} \right)^2}{2 \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right)} \right] \propto \\
&\propto \exp \left[ -\frac{1}{2} \left( \frac{1}{\sigma^2} + \frac{1}{\tau^2} \right) \left( \theta_0 - \frac{\frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}} \right)^2 \right] \propto \\
&\propto \frac{1}{\sqrt{2\pi \frac{1}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}}} \exp \left[ -\frac{1}{2 \frac{1}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}} \left( \theta_0 - \frac{\frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}} \right)^2 \right]
\end{aligned}$$

So...

$$\theta_0|\theta_1, \theta_2 \sim N \left( \frac{\frac{\theta_1}{\sigma^2} + \frac{\theta_2}{\tau^2}}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}}, \frac{1}{\frac{1}{\sigma^2} + \frac{1}{\tau^2}} \right).$$

This is a lot more intuitive in terms of **precisions**  $h_1 = 1/\sigma^2$  and  $h_2 = 1/\tau^2$ .

The precision is the inverse of the variance: more noise, less precision.

Written in terms of precisions, the above result is

$$\theta_0|\theta_1, \theta_2 \sim N\left(\frac{h_1\theta_1 + h_2\theta_2}{h_1 + h_2}, \frac{1}{h_1 + h_2}\right).$$

So, what happened?

- We got a signal  $\theta_1$  with precision  $h_1$ .
- We got a signal  $\theta_2$  with precision  $h_2$ .
- Then, **our estimate** of  $\theta_0$  is a **weighted average** of  $\theta_1$  and  $\theta_2$ , where we weight the signals according to their precision.
- And **the precision** of our estimate is **the sum** of the precisions of the signals.



This generalizes to  $n$  signals, in which case we get DeGroot's formula

$$\theta_0 | \theta_1, \dots, \theta_n \sim N \left( \frac{h_1 \theta_1 + \dots + h_n \theta_n}{h_1 + \dots + h_n}, \frac{1}{h_1 + \dots + h_n} \right).$$

This formula is key in models of career concerns (Holmström 1999, ReStud, originally published in an obscure book in 1982). Those models are the basis of models in electoral accountability (see Scott Ashworth's papers). So, it's a useful formula.

The normal distribution is very special in this regard. In general distributions do not behave as nicely.

Look up “conjugate prior” if you want to know what happens with other distributions.

# CORRELATED SIGNALS

There is a state  $\theta \in \mathbb{R}$  that nobody knows.

Two leaders, 1 and 2, communicate signals  $s_{1i}$  and  $s_{2i}$  to a group of agents indexed by  $i$ . Each agent  $i$  knows her signals  $s_{1i}$  and  $s_{2i}$  but doesn't know the signals of other agents.

For every leader  $j$  and agent  $i$  we have

$$s_{ji} = \theta + \eta_j + \epsilon_{ij},$$

with  $\eta_j \sim N(0, \sigma_j^2)$  and  $\epsilon_{ij} \sim N(0, \tau_j^2)$ , all jointly independent.

Two questions:

- Calculate  $\mathbb{E}_i[\theta \mid s_{1i}, s_{2i}]$ .
- Calculate  $\mathbb{E}_i[s_{1k} \mid s_{1i}, s_{2i}]$  for  $k \neq i$ .